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SOME TOPICS IN STATISTICAL INFORMATION
THEORY

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13. ABSTRACT <p>Attention is focused on informational properties of sub-sigma-algebras of the fundamental probability space in contrast to the discussion in <u>Information Theory and Statistics</u> (Wiley 1959, Dover Publications Inc. 1968) where attention is devoted to informational properties of statistics, that is, random variables. Properties of relative conditional expectations and results on separable sigma-algebras are stated for later use. The discrimination information in a sigma-algebra (sub-sigma-algebra) is defined and various properties developed, in particular the relation between information and sufficiency. The integral representative theorem for discrimination information is derived by methods believed to be more inherently information-theoretic than others that have been presented. Monotonic properties of conditional discrimination information are derived with respect to either the conditioning sub-sigma-algebra or the conditioned sub-sigma-algebra.</p>			

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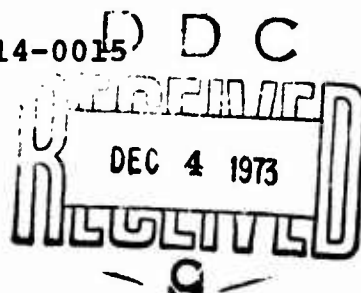
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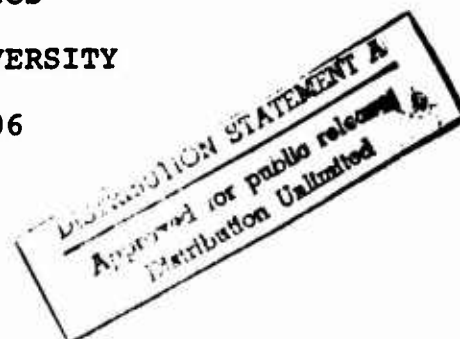
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SOME TOPICS IN STATISTICAL INFORMATION THEORY

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S. Kullback

Summary

Attention is focused on informational properties of sub-sigma-algebras of the fundamental probability space in contrast to the discussion in Information Theory and Statistics where attention is devoted to informational properties of statistics that is, random variables. In particular, the integral representation theorem for discrimination information is derived by methods believed to be more inherently information-theoretic than others that have been presented. Monotonic properties of conditional discrimination information are derived.

0. Preliminaries. In [11] attention was devoted to informational properties of statistics, that is, random variables. In this exposition however, the discussion deals with informational properties of sub-sigma-algebras of the fundamental probability space. In particular, we shall present a proof of the integral representation theorem of discrimination information which is believed to be more information-theoretic in approach than other

derivations of this basic result.

We present here certain notations, lemmas, and results on separable sigma-algebras which we shall use in this exposition.

We shall operate in the probability space (Ω, A, P) . Let $Z_t(\omega)$ and $Z(\omega)$ be non-negative random variables such that

$$(0.1) \quad \mu_t(A) = \int_A Z_t(\omega) dP(\omega), \quad \mu(A) = \int_A Z(\omega) dP(\omega), \quad A \in A$$

are probability measures. We also write (0.1) in the Radon-Nikodym differential formalism as

$$(0.2) \quad d\mu_t = Z_t dP, \quad d\mu = Z dP.$$

$Z_t(\omega)$ and $Z(\omega)$ may be considered as generalized probability densities. If we assume that μ_t is absolutely continuous with respect to μ , that is, $\mu_t \ll \mu$, then

$$(0.3) \quad d\mu_t = W_t d\mu = W_t Z dP = Z_t dP, \quad W_t = Z_t/Z \text{ a.s.}$$

so that W_t is a likelihood ratio. We shall also require sequences of the generalized densities, corresponding probability measures, and likelihood ratios, that is,

$$d\mu_{t_n} = Z_{t_n} dP, \quad d\mu_n = Z_n dP, \quad n = 1, 2, \dots$$

$$(0.4) \quad d\mu_{t_n} = W_{t_n} d\mu_n = W_{t_n} Z_n dP = Z_{t_n} dP, \quad W_{t_n} = Z_{t_n}/Z_n \text{ a.s.}$$

We shall have occasion to deal with the properties of relative conditional expectations as described in [13, p. 344].

Let B be a sub-sigma-algebra of the sigma-algebra A , that is, $B \subset A$. Corresponding to (0.2) and (0.3)

$$(0.5) \quad d\mu_B = E^B Z_t dP_B, \quad d\mu_{tB} = E^B Z_t dP_B, \quad d\mu_{tB} = E_Z^B W_t d\mu_B \\ E_Z^B W_t = E_Z^B \frac{Z_t}{Z} = (E^B Z \cdot \frac{Z_t}{Z}) / E^B Z = E^B Z_t / E^B Z,$$

where P_B is the restriction of P to B defined by $P_B(B) = P(B)$,

$B \in \mathcal{B}$, and

$$(0.6) \quad \mu_B(B) = \int_B Z dP = \int_B (E^B Z) dP_B, \quad B \in \mathcal{B}$$

$$(0.7) \quad \int_B (E_Z^B X) d\mu_B = \int_B X d\mu, \quad B \in \mathcal{B}$$

$$(0.8) \quad E^B ZX = E^B Z \cdot E_Z^B X.$$

We shall need two results from probability theory (see for example, [13, p. 140, prob. 16, 17] which we state as lemmas.

Lemma 0.1. $\int |Z_{t_n} - Z_t| dP \rightarrow 0$, resp.

$\int |Z_n - Z| dP \rightarrow 0$ as $n \rightarrow \infty$, if and only if

$\int_A Z_{t_n} dP \rightarrow \int_A Z_t dP$, resp. $\int_A Z_n dP \rightarrow \int_A Z dP$ as $n \rightarrow \infty$

uniformly in $A \in \mathcal{A}$.

Lemma 0.2. If $Z_{t_n} \xrightarrow{P} Z_t$ resp. $Z_n \xrightarrow{P} Z$, then

$\int_A Z_{t_n} dP \rightarrow \int_A Z_t dP$ resp. $\int_A Z_n dP \rightarrow \int_A Z dP$ as $n \rightarrow \infty$ uniformly in

$A \in \mathcal{A}$. The convergence in probability may be replaced by almost sure convergence.

Note that Lemmas 0.1 and 0.2 provide the chain of implications

$$(0.9) \quad \mu_n(A) \rightarrow \mu(A), \text{ uniformly in } A \in \mathcal{A} \xrightarrow{L_1} Z \Rightarrow$$

$$\Rightarrow Z_n \xrightarrow{P} Z \Rightarrow \mu_n(A) \rightarrow \mu(A), \text{ uniformly in } A \in \mathcal{A} \text{ and a similar}$$

one with the subscript t , where

$$(0.10) \quad Z_n \xrightarrow{L_1} Z \Leftrightarrow \int |Z_n - Z| dP \rightarrow 0.$$

We assemble here certain results on separable sigma-algebras which we shall need [13].

(0.11) A separable sigma-algebra is a sigma-algebra that is generated by (is minimal over) a countable class of sets.

(0.12) The minimal sigma-algebra over the union of a countable class of separable sigma-algebras is also a separable sigma-algebra.

(0.13) The Borel sigma-algebra on the real line is separable.

(0.14) The inverse image of a separable sigma-algebra by a measurable transformation is a separable sigma-algebra.

(0.15) The sub-sigma-algebra induced by a random variable or a countable class of random variables is a separable sigma-algebra.

(0.16) A finite (countable) partition of a space Ω is a finite (countable) sequence of sets A_i such that

$$\bigcup A_i = \Omega \quad A_i \cap A_j = \emptyset \quad i \neq j.$$

If A is a sigma-algebra of subsets of Ω then the partition is measurable A if $A_i \in A$ for all i . Let $E = \{E_i\}$ be an A -measurable partition. The A -measurable partition $\mathcal{D} = \{D_i\}$ is said to be a subpartition of E or finer than the partition E if each $D_i \in \mathcal{D}$ is such that $D_i \subseteq E_i \in E$ and we denote this by $\mathcal{D} < E$ or $E > \mathcal{D}$.

(0.17) A sequence of partitions $\{E^n\}$ is said to be regular if each E^n is a finite partition and

$$E^1 \supset E^2 \supset E^3 \supset \dots$$

(0.18) Let \underline{E}^n denote the finite algebra generated by the partition E^n , then corresponding to (0.17)

$$\underline{E}^1 \subset \underline{E}^2 \subset \underline{E}^3 \subset \dots$$

and $\bigcup_{n=1}^{\infty} \underline{E}^n$ is an algebra. Let \underline{E} be the minimal sigma-algebra over $\bigcup_{n=1}^{\infty} \underline{E}^n$, then \underline{E} is said to be generated by the regular sequence of partitions (0.17)

(0.19) It is clear that if a sigma-algebra is generated by a regular sequence of partitions then it is separable. Conversely, every separable sigma-algebra Λ can be generated by a regular sequence of partitions.

1. Introduction. The result in Corollary 3.2, page 16 [11] suggests that the discrimination information in the sub-sigma-algebra $B \subset A$ generated by the partition $\{B_i\}$, $i = 1, 2, \dots, n$, $B_i \in A$, $\sum_{i=1}^n B_i = \Omega$, be defined by (we shall use natural logarithms)

$$(1.1) \quad I(B; \mu_t, \mu) = \sum_{i=1}^n \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)}, \quad B_i \in B \subset A.$$

Because of the convexity of the function $x \log \frac{x}{y}$ for non-negative x and y , and additivity of the measures for disjoint sets, for $A_1, A_2 \in A$, $A_1 \cap A_2 = \emptyset$,

$$\begin{aligned} (1.2) \quad & \mu_t(A_1) \ln \frac{\mu_t(A_1)}{\mu(A_1)} + \mu_t(A_2) \ln \frac{\mu_t(A_2)}{\mu(A_2)} \\ & \geq (\mu_t(A_1) + \mu_t(A_2)) \ln \frac{\mu_t(A_1) + \mu_t(A_2)}{\mu(A_1) + \mu(A_2)} \\ & = \mu_t(A_1 + A_2) \ln \frac{\mu_t(A_1 + A_2)}{\mu(A_1 + A_2)}. \end{aligned}$$

The property in (1.2) suggests that the discrimination information in A be defined by (cf. [1], [2], [3], [6], [7], [8], [10], [14], [15])

$$(1.3) \quad \bar{I}(A; \mu_t, \mu) = \sup_{A_i \in A} \sum \mu_t(A_i) \ln \frac{\mu_t(A_i)}{\mu(A_i)}$$

where the sup is taken over all possible A -measurable finite partitions of Ω . For convenience hereafter we shall omit the μ_t and μ in $I(B; \mu_t, \mu)$ and $I(A; \mu_t, \mu)$ unless needed for clarification.

If μ_t is not absolutely continuous with respect to μ ,

that is, there exists an $A \in \mathcal{A}$ such that $\mu(A) = 0$, $\mu_t(A) \neq 0$ then $\bar{I}(A) = \infty$. Accordingly we shall assume that $\mu_t \ll \mu$. Note that $\bar{I}(A)$ may be infinite in this case also. $\mu_t \ll \mu$ is a necessary condition for $\bar{I}(A) < \infty$. (See [11, pp. 5, 6. Prob 5.7, p. 10])

It also seems intuitively reasonable to have defined the discrimination information in A by (cf. [11, p. 5])

$$(1.4) \quad I(A) = \int Z_t(\omega) \ln \frac{Z_t(\omega)}{Z(\omega)} dP(\omega).$$

The integral representation (1.4) may also be written as

$$(1.5) \quad I(A) = \int Z_t \ln \frac{Z_t}{Z} dP = \int W_t \ln W_t d\mu = \int (\ln W_t) d\mu_t = \int d\mu_t \ln \frac{d\mu_t}{d\mu}.$$

Using the additivity of the integral and Jensen's inequality for convex functions [11, p. 16]

$$(1.6) \quad \int W_t \ln W_t d\mu = \sum_{i=1}^n \int_{A_i} W_t \ln W_t d\mu, \quad A_i \in \mathcal{A}, \quad \sum_{i=1}^n A_i = \Omega,$$

$$(1.7) \quad \int_{A_i} W_t \ln W_t d\mu \geq \int_{A_i} W_t d\mu \ln \frac{\int_{A_i} W_t d\mu}{\int_{A_i} d\mu} = \mu_t(A_i) \ln \frac{\mu_t(A_i)}{\mu(A_i)},$$

$$(1.8) \quad \int W_t \ln W_t d\mu \geq \sum_{i=1}^n \mu_t(A_i) \ln \frac{\mu_t(A_i)}{\mu(A_i)},$$

hence

$$(1.9) \quad I(A) = \int W_t \ln W_t d\mu \geq \sup_{A_i \in \mathcal{A}} \sum \mu_t(A_i) \ln \frac{\mu_t(A_i)}{\mu(A_i)} = \bar{I}(A),$$

where the sup is taken over all possible \mathcal{A} -measurable finite partitions of Ω . A proof of the reverse of the inequality in (1.9) may be obtained following the method in [15, pp. 24-25]

or that in [8]. We shall not pursue this matter any further at this point but state the integral representation theorem

Theorem 1.1. $\bar{I}(A) = I(A)$, that is (1.3) and (1.5)

define the same value of the discrimination information.

Proofs of this theorem have involved martingale theory (e.g. [10]) or the use in [8] of the convexity property in conjunction with the Darboux-Young approach to the integral [13, p. 143 Ex 29].

The proof to be presented later in this exposition is believed to be intrinsically more information-theoretic in nature. For other approaches see [2],[7],[14],[15].

Note that if in the probability space (Ω, A, P) , A is generated by a finite partition $\{A_i\}$, then (0.1) is

$$(1.10) \quad \mu_t(A_i) = Z_t(\omega)P(A_i), \quad \mu(A_i) = Z(\omega)P(A_i), \quad \omega \in A_i$$

and both (1.3) and (1.5) yield

$$(1.11) \quad \bar{I}(A) = \sum_{i=1}^n \mu_t(A_i) \ln \frac{\mu_t(A_i)}{\mu(A_i)} = I(A).$$

We remark that instead of starting with the probability space (Ω, A, P) we could have started with the measure space (Ω, A, λ) where λ is a sigma-finite measure on A . We assume the existence of the non-negative A -measurable function $X(\omega)$ such that

$$(1.12) \quad P(A) = \int_A X(\omega) d\lambda(\omega), \quad A \in A$$

is a probability measure on A . We then have (see [11, p. 5])

$$(1.13) \quad d\mu_t = Z_t dP = Z_t X d\lambda = f_t d\lambda, \quad d\mu = Z dP = Z X d\lambda = f d\lambda[\lambda]$$

$$(1.14) \quad \int Z_t \ln \frac{Z_t}{Z} dP = \int Z_t \ln \frac{Z_t}{Z} X d\lambda = \int f_t \ln \frac{f_t}{f} d\lambda.$$

2. Discrimination Information in a Sub-sigma-algebra. The following discussion essentially extends some of the presentation in [11, pages 1-78]. We shall use the notation in [13] and in particular properties developed in [13, Chapter VII, Conditioning, pp. 337 ff]. In particular we shall use the fact that Jensen's inequality holds a.s. for conditional expectation also. Let \mathcal{B} be a sub-sigma-algebra of the sigma-algebra \mathcal{A} . The basic inequality is that if g is a convex function and EX is finite, then

$$(2.1) \quad E\{g(X)\} \geq g(EX).$$

In the conditional form, we have

$$(2.2) \quad E^{\mathcal{B}}\{g(X)\} \geq g(E^{\mathcal{B}}X) \text{ a.s.}$$

Using the definition of conditional expectation, (2.2) and (1.4),

$$(2.3) \quad I(\mathcal{A}) = \int Z_t \ln \frac{Z_t}{Z} dP = \int E^{\mathcal{B}}(Z_t \ln \frac{Z_t}{Z}) dP_{\mathcal{B}} \\ \geq \int E^{\mathcal{B}}Z_t \ln \frac{E^{\mathcal{B}}Z_t}{E^{\mathcal{B}}Z} dP_{\mathcal{B}}.$$

We define the right-hand side of (2.3) as the discrimination information in the sub-sigma-algebra \mathcal{B} and note that, using

(0.5),

$$(2.4) \quad I(\mathcal{B}) = \int E^{\mathcal{B}}Z_t \ln \frac{E^{\mathcal{B}}Z_t}{E^{\mathcal{B}}Z} dP_{\mathcal{B}} = \int E_Z^{\mathcal{B}}W_t \ln E_Z^{\mathcal{B}}W_t d\mu_{\mathcal{B}} \\ = \int (\ln E_Z^{\mathcal{B}}W_t) d\mu_{t, \mathcal{B}} = \int d\mu_{t, \mathcal{B}} \ln \frac{d\mu_{t, \mathcal{B}}}{d\mu_{\mathcal{B}}} \\ = \sup \int \mu_t(B_k) \ln \frac{\mu_t(B_k)}{\mu(B_k)}$$

where the Sup is taken over all finite \mathcal{B} -measurable partitions of Ω (see (2.15) and section 1). We can now state:

Theorem 2.1. If \mathcal{B} is a sub-sigma-algebra of \mathcal{A} then

$$(2.5) \quad I(\mathcal{A}) \geq I(\mathcal{B}).$$

Note that the coarsest possible sub-sigma-algebra is that generated by (\emptyset, Ω) and denoting it by \mathcal{B}_0 , $E^{\mathcal{B}_0} Z_t = EZ_t = 1$, $E^{\mathcal{B}_0} Z = EZ = 1$ and

$$(2.6) \quad I(\mathcal{B}_0) = 0.$$

Theorem 2.2. $I(\mathcal{A}) \geq 0$ with equality if and only if $W_t = Z_t/Z = 1$ a.s. (See [11, Theorem 3.1, p. 14]).

From (1.5) and (2.4), using the result that

$$(2.7) \quad \int W_t \ln E_Z^{\mathcal{B}} W_t d\mu = \int E_Z^{\mathcal{B}} (W_t \ln E_Z^{\mathcal{B}} W_t) d\mu_{\mathcal{B}} = \int E_Z^{\mathcal{B}} W_t \ln E_Z^{\mathcal{B}} W_t d\mu_{\mathcal{B}}$$

we have

$$\begin{aligned} (2.8) \quad I(\mathcal{A}) - I(\mathcal{B}) &= \int W_t \ln W_t d\mu - \int E_Z^{\mathcal{B}} W_t \ln E_Z^{\mathcal{B}} W_t d\mu_{\mathcal{B}} \\ &= \int W_t \ln (W_t / E_Z^{\mathcal{B}} W_t) d\mu \\ &= \int \frac{Z_t}{Z} \ln \frac{Z_t/Z}{E_Z^{\mathcal{B}} Z_t / E_Z^{\mathcal{B}} Z} d\mu. \end{aligned}$$

Since

$$\begin{aligned} \int W_t d\mu &= \int \frac{Z_t}{Z} d\mu = \int Z_t dP = 1, \\ (2.9) \quad \int E_Z^{\mathcal{B}} W_t d\mu_{\mathcal{B}} &= \int (E_Z^{\mathcal{B}} Z_t / E_Z^{\mathcal{B}} Z) d\mu_{\mathcal{B}} \\ &= \int E_Z^{\mathcal{B}} Z_t dP_{\mathcal{B}} = \int Z_t dP = 1, \end{aligned}$$

the right-hand side of (2.8) is a discrimination information value, and as such non-negative. Let us define, using (2.8)

above, and (0.2) and (0.5)

$$(2.10) \quad I(A|B) = \int W_t \ln (W_t / E_Z^B W_t) d\mu = \int Z_t \ln \frac{Z_t / E_Z^B Z_t}{Z / E_Z^B Z} dP = \\ = \int d\mu_t \ln \frac{d\mu_t / d\mu_B}{d\mu / d\mu_B} \quad ,$$

that is, $I(A|B)$ is the conditional discrimination information in A given B and hence

Theorem 2.3. If B is a sub-sigma-algebra of A

$$(2.11) \quad I(A) = I(B) + I(A|B).$$

Theorem 2.4. If B is a sub-sigma-algebra of A , then

$$(2.12) \quad I(A) = I(B)$$

if and only if $I(A|B) = 0$, that is, if and only if

$$(2.13) \quad Z_t / Z = E_Z^B Z_t / E_Z^B Z. \quad \text{a.s.}$$

Proof. Apply Theorem 2.2 and (2.10).

If B is a sub-sigma-algebra of A and satisfies Theorem 2.4, then we say that B is a sufficient sub-sigma-algebra for A . (See [13, p. 346], [11, pp. 18-22].)

If B is generated by a finite partition, then on a non-null atom $B \in B$, the conditional expected value $E_Z^B X$ is a constant and its value is

$$(2.14) \quad E_Z^B X = \frac{1}{P(B)} \int_B X dP, \quad \omega \in B.$$

Thus

$$(2.15) \quad I(B) = \int E_Z^B Z_t \ln \frac{E_Z^B Z_t}{E_Z^B Z} dP_B = \int \int_{B_i} E_Z^B Z_t \ln \frac{E_Z^B Z_t}{E_Z^B Z} dP_B \\ = \sum_i P(B_i) \cdot \frac{1}{P(B_i)} \int_{B_i} Z_t dP \ln \frac{\int_{B_i} Z_t dP}{\int_{B_i} Z dP} = \sum_i \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)} .$$

Note that in this case

$$(2.16) \quad \frac{Z}{E^B Z} = Z / \frac{1}{P(B)} \quad \int_B Z dP = Z / (\mu(B) / P(B)), \quad \omega \in B$$

and

$$\begin{aligned} (2.17) \quad I(A|B) &= \int Z_t \ln \frac{Z_t / E^B Z_t}{Z / E^B Z} dP \\ &= \sum_i \int_{B_i} Z_t \ln \frac{Z_t / E^B Z_t}{Z / E^B Z} dP \\ &= \sum_i \int_{B_i} Z_t \ln \frac{Z_t P(B_i) \mu(B_i)}{\mu_t(B_i) Z P(B_i)} dP \\ &= \sum_i \mu_t(B_i) \int_{B_i} \frac{Z_t}{\mu_t(B_i)} \ln \frac{Z_t / \mu_t(B_i)}{Z / \mu(B_i)} dP \\ &= \sum_i \mu_t(B_i) I(A|B_i) \end{aligned}$$

where

$$\begin{aligned} (2.18) \quad I(A|B_i) &= \int_{B_i} \frac{Z_t}{\mu_t(B_i)} \ln \frac{Z_t / \mu_t(B_i)}{Z / \mu(B_i)} dP \\ &= \int_{B_i} \frac{d\mu_t}{\mu_t(B_i)} \ln \frac{d\mu_t / \mu_t(B_i)}{d\mu / \mu(B_i)} \\ &= \frac{1}{\mu_t(B_i)} \int_{B_i} Z_t \ln \frac{Z_t}{Z} dP - \ln \frac{\mu_t(B_i)}{\mu(B_i)} \\ &= \frac{1}{\mu_t(B_i)} \int_{B_i} W_t \ln W_t d\mu - \ln \frac{\mu_t(B_i)}{\mu(B_i)}. \end{aligned}$$

Note that $I(A|B_i)$ is a discrimination information, and

$$(2.19) \quad I(A|B_i) \geq 0$$

with equality if and only if

$$(2.20) \quad Z_t/Z = \mu_t(B_1)/\mu(B_1), \quad \omega \in B_1,$$

(see [11, Corollary 3.1, p. 15].)

To obtain another representation for $I(A|B)$ when B is a separable sub-sigma-algebra of A we proceed as follows. It follows from (2.18) that

$$\begin{aligned} (2.21) \quad \sum_i \mu_t(B_i) I(A|B_i) &= \sum_i \int_{B_i} Z_t \ln \frac{Z_t}{Z} dP - \sum_i \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)} \\ &= \int Z_t \ln \frac{Z_t}{Z} dP - \sum_i \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)} \\ &= I(A) - \sum_i \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)}. \end{aligned}$$

Now let $\Pi(B)$ denote the class of finite B -measurable partitions of Ω , then (see [8])

$$\begin{aligned} (2.22) \quad \inf_{\Pi(B)} \sum_i \mu_t(B_i) I(A|B_i) &= \inf_{\Pi(B)} \{ I(A) - \sum_i \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)} \} \\ &= I(A) - \sup_{\Pi(B)} \sum_i \mu_t(B_i) \ln \frac{\mu_t(B_i)}{\mu(B_i)} \\ &= I(A) - I(B) = I(A|B). \end{aligned}$$

From (2.22) we note that $I(B|B) = 0$.

Let W_t be defined as in (0.3) and denote by \mathcal{B}_{W_t} the class of sets $W_t^{-1}(B)$ where B ranges over linear Borel sets, that is

$$(2.23) \quad W_t^{-1}(B) = \{\omega: W_t(\omega) \in B\}.$$

\mathcal{B}_{W_t} is the minimal sigma-algebra with respect to which W_t is measurable and \mathcal{B}_{W_t} is a separable sub-sigma-algebra of \mathcal{A} . For convenience we shall hereafter denote \mathcal{B}_{W_t} by \mathcal{B}_t . We shall now show that \mathcal{B}_t is a sufficient sub-sigma-algebra for \mathcal{A} .

Theorem 2.5. \mathcal{B}_t is a sufficient sub-sigma-algebra for \mathcal{A} , that is, $I(\mathcal{A}) = I(\mathcal{B}_t)$.

Proof.

$$\begin{aligned} (2.24) \quad I(\mathcal{A}) &= \int W_t \ln W_t d\mu = \int E_Z^{\mathcal{B}_t}(W_t \ln W_t) d\mu_{\mathcal{B}_t} \\ &\geq \int E_Z^{\mathcal{B}_t} W_t \ln E_Z^{\mathcal{B}_t} W_t d\mu_{\mathcal{B}_t} = I(\mathcal{B}_t) = \int W_t \ln W_t d\mu = I(\mathcal{A}) \end{aligned}$$

where we have used the fact that since W_t is measurable \mathcal{B}_t then $E_Z^{\mathcal{B}_t} W_t = W_t [\mu]$. Note that using (0.5)

$$E^{\mathcal{B}_t} Z_t / E^{\mathcal{B}_t} Z = Z_t / Z,$$

which is the necessary and sufficient condition (2.13) for Theorem 2.4.

3. An Information-theoretic Approach. We shall now consider an approach to the integral representation which is believed to be of interest in that it is intrinsically information-theoretic in nature.

Suppose that there is an increasing sequence of sub-sigma-algebras of Λ such that

$$(3.0) \quad B_0 \subset B_1 \subset B_2 \subset \dots \subset B_n \subset \dots \subset B \subset \Lambda$$

where B is the minimal sigma-algebra containing $\bigcup_n B_n$,

usually denoted by $B_n \hat{\bigcup} B$, or $B = \bigvee_n B_n$.

Consider

$$(3.1) \quad \int E^B Z \ln \frac{E^B Z}{E^n Z} dP = \int E^B Z \ln \frac{E^{B_{n+1}} Z}{E^n Z} dP + \int E^B Z \ln \frac{E^B Z}{E^{B_{n+1}} Z} dP$$

so that since

$$(3.2) \quad \begin{aligned} \int E^B Z \ln \frac{E^{B_{n+1}} Z}{E^n Z} dP &= \int E^{B_{n+1}} (E^B Z \ln \frac{E^{B_{n+1}} Z}{E^n Z}) dP \\ &= \int E^{B_{n+1}} Z \ln \frac{E^{B_{n+1}} Z}{E^n Z} dP \geq 0 \end{aligned}$$

(recall that $E^{B_{n+1}} (E^B Z) = E^{B_{n+1}} Z$ and $E^{B_{n+1}} (E^n Z) = E^n Z$) we may write (3.1) as

$$(3.3) \quad \begin{aligned} \int E^B Z \ln \frac{E^B Z}{E^n Z} dP &= \int E^{B_{n+1}} Z \ln \frac{E^{B_{n+1}} Z}{E^n Z} dP + \\ &= \int E^B Z \ln \frac{E^B Z}{E^{B_{n+1}} Z} dP. \end{aligned}$$

Hence using the notation

$$(3.4) \quad I(B:B_n) = \int E^{B_n} Z \ln \frac{E^B Z}{E^{B_n} Z} dP$$

we have the chain of inequalities

$$(3.5) \quad I(B:B_0) \geq I(B:B_1) \geq \dots \geq I(B:B_n) \geq I(B:B_{n+1}) \geq \dots \geq 0$$

where it is seen from (3.3) that $I(B:B_n) = I(B:B_{n+1}) =$

$\Leftrightarrow I(B_{n+1}:B_n) = 0$. Since the monotonically decreasing

sequence (3.5) is bounded below it converges and hence

$$(3.6) \quad I(B:B_n) - I(B:B_{n+1}) = \int E^{B_{n+1}} Z \ln \frac{E^{B_{n+1}} Z}{E^{B_n} Z} dP \rightarrow 0, \quad n \rightarrow \infty.$$

Hence using the result in [12], (3.6) implies

$$(3.7) \quad \int |E^{B_{n+1}} Z - E^{B_n} Z| dP \rightarrow 0, \quad n \rightarrow \infty.$$

Indeed, by considering $I(B:B_{n+m}) - I(B:B_n)$ we can also get

$$(3.8) \quad \int |E^{B_{n+m}} Z - E^{B_n} Z| dP \rightarrow 0 \quad m, n \rightarrow \infty,$$

hence the sequence $E^{B_n} Z$ is L_1 fundamental and there exists an $X \in L_1$ such that

$$(3.9) \quad \int |E^{B_n} Z - X| dP \rightarrow 0, \quad n \rightarrow \infty,$$

and

$$(3.10) \quad \int E^{B_n} Z dP \rightarrow \int X dP, \quad n \rightarrow \infty,$$

that is

$$(3.11) \quad \int X dP = 1$$

since $\int E^{B_n} Z dP = 1$ for all n (see [13, p. 157, 161]).

Without restricting the generality we may take X to be \mathcal{B} -measurable by the following argument. Using the result in [13, p. 348] that

$$X_n \xrightarrow{L_1} X \Rightarrow E^{\mathcal{B}} X_n \xrightarrow{L_1} E^{\mathcal{B}} X$$

we have

$$\int |E^{\mathcal{B}_n} Z - X| dP \rightarrow 0, n \rightarrow \infty \Rightarrow \int |E^{\mathcal{B}} E^{\mathcal{B}_n} Z - E^{\mathcal{B}} X| dP \rightarrow 0,$$

$$n \rightarrow \infty$$

$$\Rightarrow \int |E^{\mathcal{B}_n} Z - E^{\mathcal{B}} X| dP \rightarrow 0, n \rightarrow \infty$$

since $E^{\mathcal{B}} E^{\mathcal{B}_n} Z = E^{\mathcal{B}_n} Z$. If X is \mathcal{B} -measurable $E^{\mathcal{B}} X = X$ a.s.

We shall now show that $X = E^{\mathcal{B}} Z$ a.s. Applying lemma 0.1 to (3.9) we have

$$(3.12) \quad \int_A E^{B_n} Z dP \rightarrow \int_A X dP.$$

If $A \in \mathcal{B}_k$, then for $n \geq k$

$$(3.13) \quad \int_A E^{B_n} Z dP = \int_A E^B Z dP$$

so that

$$(3.14) \quad \int_A X dP = \int_A E^B Z dP.$$

Since (3.14) is true for $A \in \mathcal{B}_k$, it is true for $A \in \mathcal{U}\mathcal{B}_k$.

Thus the probability measures defined by the integrals in (3.14) are identical on the field $\mathcal{U}\mathcal{B}_k$ and hence by the extension theorem [13, p. 87] (3.14) holds for $A \in \mathcal{B}$. The Radon-Nikodym theorem then yields from (3.14) that

$$(3.15) \quad X = E^B Z [P_B].$$

Since

$$(3.16) \quad \int |E^{B_n} Z - E^B Z| dP \leq \int |E^{B_n} Z - X| dP + \int |X - E^B Z| dP$$

we see from (3.9) and (3.15) that (cf. [4, pp. 319, 331])

$$(3.17) \quad \int |E^{B_n} Z - E^B Z| dP \rightarrow 0, \quad n \rightarrow \infty.$$

Since, using theorem 2.2

$$(3.18) \quad \int E^B Z \ln \frac{E^B Z}{X} dP = 0$$

if and only if $X = E^B Z$ a.s., then in view of (3.15) we conjecture that the sequence (3.5) has the limit zero, that is,

$$(3.19) \quad \lim_{n \rightarrow \infty} I(\mathcal{B}; \mathcal{B}_n) = 0$$

We show that (3.19) is true as follows. We write

$$\begin{aligned} I(B:B_n) &= \int E^B Z \ln \frac{E^B Z}{E^{B_n} Z} dP = \int E^B Z \ln E^B Z dP - \int E^B Z \ln E^{B_n} Z dP \\ &= \int E^B Z \ln E^B Z dP - \int E^{B_n} Z \ln E^{B_n} Z dP. \end{aligned}$$

We have shown that $E^{B_n} Z \xrightarrow{L^1} E^B Z$ which implies that $E^{B_n} Z \xrightarrow{P} E^B Z$.

The convergence in probability implies that there exists a sequence $\{n_k\}$ of integers increasing to infinity such that [13, p. 151]

$$E^{B_{n_k}} Z \xrightarrow{a.s.} E^B Z.$$

Since the convex function $E^{B_n} Z \ln E^{B_n} Z \geq -1/e$, the Fatou-Lebesgue Theorem [13, p. 152] yields

$$(3.20) \quad \liminf \int E^{B_{n_k}} Z \ln E^{B_{n_k}} Z dP \geq \int E^B Z \ln E^B Z dP.$$

But the convexity, Jensen's inequality, and the smoothing property of conditioning [13, p. 348, 351] lead to

$$(3.21) \quad \int E^B Z \ln E^B Z dP \geq \int E^{B_n} Z \ln E^{B_n} Z dP \geq \int E^{B_m} Z \ln E^{B_m} Z dP$$

for all n, m such that $n > m$.

From the monotonic property in (3.21) combined with

(3.20) we conclude that

$$\lim \int E^{B_{n_k}} Z \ln E^{B_{n_k}} Z dP = \int E^B Z \ln E^B Z dP$$

or

$$\lim I(B:B_{n_k}) = 0.$$

Since $I(B:B_n)$ converges, it must converge to the same limit, that is, we have (3.19).

Note that if we interpret $I(B:B_n)$ as a measure of the closeness of the approximation to the measures over B by the measures over B_n , the sequence (3.5) implies that the approximation gets better as n gets larger and in the limit the approximation is exact to within sets of measure zero.

A similar argument of course follows for Z_t . If we start with

$$(3.22) \quad \int_{E_Z^B W_t} \ln \frac{E_Z^B W_t}{E_Z^{B_n} W_t} d\mu = \int_{E_Z^B W_t} \ln \frac{E_Z^{B_{n+1}} W_t}{E_Z^{B_n} W_t} d\mu + \\ + \int_{E_Z^B W_t} \ln \frac{E_Z^B W_t}{E_Z^{B_{n+1}} W_t} d\mu$$

since

$$(3.23) \quad \int_{E_Z^B W_t} \ln \frac{E_Z^{B_{n+1}} W_t}{E_Z^{B_n} W_t} d\mu = \int_{E_Z^{B_{n+1}} W_t} \ln \frac{E_Z^{B_{n+1}} W_t}{E_Z^{B_n} W_t} d\mu$$

we can repeat the preceding argument and conclude that

$$(3.24) \quad \begin{cases} \int |E_Z^{B_n} W_t - Y_t| d\mu \rightarrow 0, & n \rightarrow \infty \\ \int Y_t d\mu = 1, \\ Y_t = E_Z^B W_t \quad [\mu_B] \end{cases}$$

$$(3.25) \quad \lim_{n \rightarrow \infty} \int E_{Z^n}^{B_{W_t}} \ln \frac{E_{Z^n}^{B_{W_t}}}{E_Z^{B_{W_t}}} d\mu = 0.$$

We shall now use these results to show that

$$(3.26) \quad I(B) = \int E_Z^{B_{W_t}} \ln E_Z^{B_{W_t}} d\mu = \lim_{n \rightarrow \infty} \int E_Z^{B_n W_t} \ln E_Z^{B_n W_t} d\mu = \\ = \lim_{n \rightarrow \infty} I(B_n).$$

Using the result in theorem 2.3 we may write

$$(3.27) \quad \begin{aligned} I(B|B_0) &= I(B_1|B_0) + I(B|B_1) \\ I(B|B_1) &= I(B_2|B_1) + I(B|B_2) \\ &\dots \dots \dots \\ I(B|B_k) &= I(B_{k+1}|B_k) + I(B|B_{k+1}) \\ &\dots \dots \dots \end{aligned}$$

or using the relations in (3.27)

$$(3.28) \quad I(B|B_0) = I(B) = \sum_{k=0}^n I(B_{k+1}|B_k) + I(B|B_{n+1})$$

where

$$(3.29) \quad I(B|B_N) = \int E_Z^{B_{W_t}} \ln \frac{E_Z^{B_{W_t}}}{E_Z^{B_N W_t}} d\mu.$$

Since the relation (3.25) applied to (3.29) implies

$$(3.30) \quad I(B|B_N) \rightarrow 0 \text{ as } N \rightarrow \infty$$

we get from (3.28) (see (2.8), (2.10))

$$(3.31) \quad I(B) = \sum_{k=0}^{\infty} I(B_{k+1}|B_k) = \sum_{k=0}^{\infty} (I(B_{k+1}) - I(B_k)) = \lim_{n \rightarrow \infty} I(B_n).$$

From (3.31) it is seen that if $I(B) < \infty$ then $I(B_{n+1}|B_n) \rightarrow 0$,

$n \rightarrow \infty$. If $I(\mathcal{B}_{n+1} | \mathcal{B}_n) \rightarrow I < \infty$ as $n \rightarrow \infty$, then by the Toeplitz lemma [13, p. 238] $\lim_{n \rightarrow \infty} \frac{1}{n} I(\mathcal{B}_n) = I$.

We are now in a position to apply (3.31) to prove the integral representation theorem.

In the preliminaries it was seen that a separable sigma-algebra can be generated by a regular sequence of partitions, and using the notation of (0.18) we have

$$(3.32) \quad \underline{E}^n \uparrow \underline{E}.$$

Accordingly, as a special case of (3.31) we have that

$$(3.33) \quad \lim_{n \rightarrow \infty} I(\underline{E}^n) = I(\underline{E}).$$

Recalling (1.11), (2.15) and Theorem 2.5, we note that (3.33) is the integral representation theorem, that is

$$(3.34) \quad I(A) = \sup \left[\mu_t(A_i) \ln \frac{\mu_t(A_i)}{\mu(A_i)} \right] = \bar{I}(A)$$

where the sup is taken over all possible \mathcal{A} -measurable finite partitions of Ω , and in particular if the \mathcal{B}_n are generated by finite partitions then (3.31) is the same as (2.4).

For application of the preceding results, particularly (3.31) to stochastic processes we state as Lemma 3.1 a result which is Theorem 1.6, page 604 of [4].

Lemma 3.1. Let \mathcal{A} be a sigma-algebra of ω sets, and let $\{x(t, \omega), t \in T\}$ be a family of ω functions measurable with

respect to Λ . Let \mathcal{B}_S be the sigma-algebra generated by $\{x(t, \omega), t \in S \subset T\}$. Suppose that T is non-denumerable. Then if $\Lambda \in \mathcal{B}_T$ there is a denumerable subset S (depending on Λ) of T , such that $\Lambda \in \mathcal{B}_S$. If $X(\omega)$ is an ω function measurable with respect to \mathcal{B}_T , there is a denumerable subset S (depending on X) such that X is measurable with respect to \mathcal{B}_S .

Now let $\{x(t, \omega), t \in T\}$ be an arbitrary system of random variables defining a stochastic process. Let \mathcal{B}_N be the sigma-algebra generated by the sub-system $\{x(t, \omega), t \in N \subset T\}$ and \mathcal{B}_T the sigma-algebra generated by the system of random variables defining the stochastic process. We can now state

Theorem 3.1.

$$(3.35) \quad I(\mathcal{B}_T) = \sup_{N \in \mathcal{N}} I(\mathcal{B}_N)$$

where \mathcal{N} is the class of all finite subsets of T .

Proof. If T is countable, then it is possible to choose finite subsets $N_1 \subset N_2 \subset \dots$ such that \mathcal{B}_T is the smallest sigma-algebra containing $\bigcup_{i=1}^{\infty} \mathcal{B}_{N_i}$ and (3.35) is then essentially

(3.31). If T is not countable we shall use Lemma 3.1. If $I(\mathcal{B}_S) = \infty$, then $I(\mathcal{B}_T) = \infty$. Suppose $I(\mathcal{B}_T) < \infty$, then

$$E_Z^{B_T} W_t = E^{B_T}_{Z_t} / E^{B_T}_Z \text{ (see (0.5)) exists and is of course}$$

measurable with respect to \mathcal{B}_T . But according to Lemma 3.1 every function measurable with respect to \mathcal{B}_T is measurable with respect to \mathcal{B}_S for at least one countable $S \subset T$ so that

$$(3.36) \quad E_Z^{\mathcal{B}_T} W_t = E_Z^{\mathcal{B}_S} W_t \quad [\mu], \text{ or } E^{\mathcal{B}_T}_{Z_t} / E^{\mathcal{B}_T}_Z = E^{\mathcal{B}_S}_{Z_t} / E^{\mathcal{B}_S}_Z \quad \text{a.s.}$$

Since (3.36) is the necessary and sufficient condition that $I(\mathcal{B}_S) = I(\mathcal{B}_T)$ we get

$$(3.37) \quad I(\mathcal{B}_T) = I(\mathcal{B}_S) = \sup_{N \in \mathcal{N}} I(\mathcal{B}_N).$$

For a similar result see [9].

4. Monotonicity. If the sigma-algebra \mathcal{B} in (3.0) is not the sigma-algebra \mathcal{A} of the probability space (Ω, \mathcal{A}, P) then again using theorem 2.3 as in (3.27) we have

$$(4.1) \quad I(\mathcal{A}) = I(\mathcal{B}) + I(\mathcal{A}|\mathcal{B})$$

$$(4.2) \quad I(\mathcal{A}) = I(\mathcal{B}_n) + I(\mathcal{A}|\mathcal{B}_n).$$

We can now derive certain limiting relations in which \mathcal{A} plays a role. From (4.2) and (3.31) we see that

$$(4.3) \quad \begin{aligned} I(\mathcal{A}) &= \lim_{n \rightarrow \infty} I(\mathcal{B}_n) + \lim_{n \rightarrow \infty} I(\mathcal{A}|\mathcal{B}_n) \\ &= I(\mathcal{B}) + \lim_{n \rightarrow \infty} I(\mathcal{A}|\mathcal{B}_n), \end{aligned}$$

hence for $I(\mathcal{A}) < \infty$

$$(4.4) \quad I(\mathcal{A}) - I(\mathcal{B}) = I(\mathcal{A}|\mathcal{B}) = \lim_{n \rightarrow \infty} I(\mathcal{A}|\mathcal{B}_n).$$

Similarly from

$$(4.5) \quad I(\mathcal{B}) = I(\mathcal{B}_n) + I(\mathcal{B}|\mathcal{B}_n)$$

$$(4.6) \quad I(B) = \lim_{n \rightarrow \infty} I(B_n) + \lim_{n \rightarrow \infty} I(B|B_n)$$

so that if $I(B) < \infty$

$$(4.7) \quad I(B) = \lim_{n \rightarrow \infty} I(B_n) \Leftrightarrow \lim_{n \rightarrow \infty} I(B|B_n) = 0.$$

Note that if $I(A) < \infty$ and $B_n \subset B_{n+1}$, then using (4.2) we have

$$(4.8) \quad I(A|B_n) - I(A|B_{n+1}) = I(B_n) - I(B_{n+1}) \geq 0.$$

Similarly, for $B_k \subset B_n \subset B_{n+1}$ we have

$$(4.9) \quad I(B_n|B_k) \geq I(B_n|B_{n+1}).$$

As a matter of fact for $B_1 \subset B_2 \subset A$ we can write

if $I(A) < \infty$ (for a related discussion see [8])

$$(4.10) \quad I(A|B_1) = I(B_2|B_1) + I(A|B_2)$$

which can be proven either directly or from the fact that

$$(4.11) \quad [I(A) - I(B_1)] = [I(B_2) - I(B_1)] + [I(A) - I(B_2)].$$

Since the information values in (4.10) are nonnegative it follows that

$$(4.12) \quad I(A|B_1) \geq I(A|B_2)$$

with equality if and only if $I(B_2|B_1) = 0$ and

$$(4.13) \quad I(A|B_1) \geq I(B_2|B_1).$$

with equality if and only if $I(A|B_2) = 0$.

Using A in place of B in (3.27) we have corresponding to (3.28)

$$(4.14) \quad I(A|B_0) = I(A) = \sum_{k=0}^n I(B_{k+1}|B_k) + I(A|B_{n+1})$$

so that using (4.4)

$$(4.15) \quad I(A) = \sum_{k=0}^{\infty} I(B_{k+1} | B_k) + I(A|B).$$

In connection with the results of this section see the axiomatic approach in [5].

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